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DOI:

[10.1017/S1352465813000556](https://doi.org/10.1017/S1352465813000556)

Document Version

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Citation for published version (APA):

Veale, D., Eshkevari, E., Kanakam, N., Ellison, N., Costa, A., & Werner, T. (2014). The Appearance Anxiety Inventory: Validation of a Process Measure in the Treatment of Body Dysmorphic Disorder. *Behavioural and Cognitive Psychotherapy*, 42(5), 605-616. <https://doi.org/10.1017/S1352465813000556>

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Running Head: The Appearance Anxiety Inventory

The Appearance Anxiety Inventory: Validation of a process measure in the treatment of Body Dysmorphic Disorder

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Abstract

Background: At present there are no measures to identify the cognitive processes and behaviours that might mediate the outcome of treatment in people with Body Dysmorphic Disorder (BDD). **Aims:** To develop and validate a process measure that can be used to assess the progress of patients throughout therapy and in research for BDD. **Method:** The psychometric properties of the Appearance Anxiety Inventory (AAI) was explored in a clinical group of participants diagnosed with BDD (Study 1) and in a non-clinical community group with high appearance concerns (Study 2). Item characteristics, reliability, and factor structure were analysed. Convergent validity with measures of related symptoms was assessed. **Results:** The AAI was found to have good test-retest reliability and convergent validity in the measurement of appearance anxiety. It was also sensitive to change during treatment. The scale was found to have a two-factor structure in the clinical group, with one factor characterised by avoidance, and a second factor comprised of threat monitoring. However, in the community sample it appeared to have a one-factor structure. **Conclusion:** The results suggest that the AAI has the psychometric properties to determine if changes in cognitive processes and behaviours can mediate the outcome following treatment in patients with BDD. This supports its potential usefulness in clinical and research settings.

Keywords: Body Dysmorphic Disorder; questionnaire; assessment

The Appearance Anxiety Inventory

One of the criticisms of cognitive behaviour therapy (CBT) is the limited evidence for the factors that mediate change (Hayes, Villate, Levin, & Hildebrant, 2011; Hofman, 2008; Longmore & Worrell, 2007; Worrell & Longmore, 2008). This observation also applies to CBT for Body Dysmorphic Disorder (BDD) where there is limited research into the factors that mediate the outcome of treatment. At present there have been two randomized controlled trials which demonstrated the benefits of CBT compared to a waiting list (Rosen, Reiter, & Orosan, 1995; Veale et al., 1996). However, there is no evidence for the processes that may mediate improvement following treatment. Such studies are challenging since they require sample sizes much larger than that for a randomized clinical trial. This would enable linear regression models to be conducted which assess the factors that predict long term response to treatment (Kraemer, Wilson, Fairburn, & Agras, 2002).

There are a number of self-report scales that measure the severity of symptoms in BDD, such as the Body Image Disturbance Questionnaire (Cash, Phillips, Santos, & Hrabosky, 2004), the Dysmorphic Concern Questionnaire (Mancuso, Knoesen, & Castle, 2010), the Cosmetic Procedure Screening Scale (Veale et al., 2011) and the impact of Body Image on the Quality of Life (Cash & Fleming, 2002; Hrabosky et al., 2009). There is also an observer rated scale: the Yale Brown Obsessive Compulsive Scale modified for BDD (Phillips et al., 1997). All these scales tend to measure predominantly the 'output' - in particular the degree of preoccupation, distress and interference in a person's life. However there are no scales that measure factors that may mediate response to treatment. Such measures could help to explore which

mediating factors are associated with subsequent changes in outcome or prognosis. It is recognised, though, that measuring the processes and output may overlap to a certain extent, and that better outcome scales focussing on distress and quality of life may therefore be required.

The items for the new scale (i.e. the Appearance Anxiety Inventory), were drawn from the literature on trans-diagnostic processes that occur in mental disorders (Harvey, Watkins, Mansell, & Shafran, 2004) and a theoretical model of the maintenance of BDD (Neziroglu, Khemlani-Patel & Veale, 2008; Veale, 2004, Veale et al, 1996). In this model, the distorted imagery and associated shame is central to the experience of the self as an aesthetic object. The responses to a distorted image and shame include various cognitive processes (e.g. self-focused attention, rumination, and comparing or mental planning of cosmetic procedures) and behaviours (e.g. appearance-checking in reflective surfaces, questioning others, avoiding people, or adopting safety-seeking behaviours such as camouflaging the perceived defect.) Furthermore, changes in distorted imagery, cognitive processes and behaviours mediate change in outcome such as preoccupation, distress and interference in life. However this model needs to be tested empirically.

The present study sought to address the need for a measure of factors that mediate change by developing and validating a new scale, the Appearance Anxiety Inventory (AAI). The AAI is a self-report measure that focuses on the cognitive processes and safety seeking behaviours that are characteristic of a response to a distorted body image and associated shame. The intention is to develop a scale that was brief, free to use, sensitive to change during treatment, and suitable for weekly assessment. The

scale aims to have the potential to determine if the theorized cognitive and behavioural processes mediate change. It also seeks to assist clinicians and clients to determine which processes to target during therapy. The psychometric properties of the AAI were therefore examined in a clinical group of patients with BDD (Study 1) and in a non-clinical community group of people with high appearance concerns (Study 2). The latter group was chosen as analogue population of people with BDD (for example they wanted to improve one or more features in their appearance and were motivated to have a cosmetic procedure). Body image concerns are often on a continuum with people with BDD on the extreme end. We would therefore expect similar processes in an analogue population of BDD but occurring a lower frequency than in BDD. It was therefore hypothesised that the BDD group should score significantly greater on the AAI than non-clinical group with high appearance concerns.

Method Study 1

Participants

A total of 139 participants with BDD participated in the first study. The median age of the participants was 28 (IQR = 12). This group consisted of 67 (48.2%) males and 72 (51.8%) females. Thirty-four (24.5%) were unemployed or on long-term sick leave, 53 (38.1%) were employed or self-employed, 2 (1.4%) were retired, 18 (12.9%) were students, and 3 (2.2%) were homemakers and 29 (21.1%) unknown. All participants completed the questionnaires before any treatment.

We report on a sub-sample of 12 participants with BDD whose AAI was repeated at weekly intervals during treatment to determine sensitivity to change. The median age

of the treatment group was 33 (IQR = 15). The treatment group consisted of 5 males and 7 females. Three (25%) were unemployed, 5 (42%) were employed or self-employed, 2 (17%) were full-time students, and 2 (17%) did not specify their occupation.

Procedure

Participants were recruited from outpatient settings. A structured diagnostic interview (SCID for DSM-IV) was used to diagnose BDD as the main problem in all participants. A SCID was not conducted for comorbidity in all the participants. The BDD group were interviewed with the Yale Brown Obsessive Compulsive Scale for BDD (BDD-YBOCS) (Phillips et al., 1997) and the Brown Assessment of Beliefs Scale (BABS) (Eisen et al., 1998). A sample (12) of the BDD group completed the AAI weekly during cognitive behaviour therapy.

Materials

1. Appearance Anxiety Inventory (AAI)

The AAI was developed through a process of iteration. Experienced clinicians and people with BDD reviewed the items and the wording was accordingly modified. It was tested in a pilot study of people with BDD before the final version was used for the study. The scale was comprised of 10 items. Each item is scored on a 5-point Likert scale ranging from zero for “not at all” to four for “all the time”. The maximum score is 40, and higher scores reflect greater frequency of a process. The total score is obtained by summing all the items. No items are reverse scored.

2. Yale Brown Obsessive Compulsive Scale modified for BDD (BDD-YBOCS)

(Phillips et al., 1997)

The BDD-YBOCS is a 12-item observer rated measure for assessing the severity of BDD symptoms that occurred during the past week. The first 5 items rate BDD-related preoccupations, and the second 5 items rate BDD-related behaviours. The BDD-YBOCS also rates one item on the degree of insight and one item on the degree of avoidance behaviour. Each item is rated on a scale 0-4 (range from 0-48). Higher summed scores indicate greater BDD symptomatology. A clinical case is likely to score 24 or more.

3. Brown Assessment of Beliefs (BABS) (Eisen et al., 1998)

The BABS was used for appearance-related beliefs that participants provided about themselves. The BABS is a 6-item observer rated measure for the degree of conviction in a belief (e.g. “I am as ugly as the Elephant man”). Items measure conviction, perception of others’ views of beliefs, explanation of differing views, fixity of ideas, attempts to disprove ideas, and insight, respectively. Each item is rated from 0-4 (from least to most severe). Total scores are summed to give a possible range of 0-24, with higher scores representing higher conviction in a belief.

4. Body Image Quality Of Life Inventory (BIQLI) (Cash & Fleming, 2002)

The BIQLI consists of 19 items on a 7-point Likert scale (with values ranging from -3 (“very negative effect”) to +3 (“very positive effect”) and measures the impact of body image concerns on a range of life domains. A lower score indicates more negative impact of body image on a person’s quality of life. Hrabosky et al. (2009) found that people with BDD scored a mean of -1.81 (standard deviation of 0.68)

compared to 0.06 (SD = .16) in female psychiatric controls and 0.4 (SD = 1.21) in male psychiatric controls.

5. The Patient Health Questionnaire (PHQ-9) (Kroenke & Spitzer, 2002).

The PHQ-9 is a self-report version of the PRIME-MD diagnostic instrument for common mental disorders. The PHQ-9 is the depression module, which scores each of the 9 DSM-IV criteria as “0” (not at all) to “3” (nearly every day). The PHQ-9 total score for the 9 items ranges from 0 to 27. A higher score represents higher depression symptomatology.

6. Generalized Anxiety Disorder assessment (GAD-7) (Spitzer, Kroenke, Williams, & Lowe, 2006).

The GAD-7 is designed primarily as a screening and severity measure for symptoms of generalized anxiety disorder. The total score of the GAD-7 is the total score of the 7 items, which ranges from 0 to 21. Scores of 5, 10, and 15 represent cut points for mild, moderate, and severe anxiety, respectively.

Statistical Analysis

Horn’s Parallel Factor Analysis (Horn, 1965) was used first to examine factorial validity of the AAI. This was computed using the factor analysis programme ‘FACTOR’ (Lorenzo-Seva & Ferrando, 2006). This method was chosen as it is more accurate than Cattell’s scree and Kaiser-Guttman methods (Wilson & Cooper, 2008; Zwick & Velicer, 1986). The internal consistency of the AAI was evaluated using Cronbach’s alpha. Total scores of the AAI were not normally distributed so that Spearman’s Rho correlations between the AAI and the BDD-YBOCS, BABS,

BIQLI, PHQ9 and GAD7 were computed to test convergent validity. Given that the total scores were not centred on a common mean we report a median and inter-quartile range as descriptors of variance. Non-parametric comparisons (Wilcoxon signed ranks tests) were used to compare pre-treatment (before session 1) and post-treatment (session 16) scores due to the small sample size not being normally distributed.

To reduce missing data from partially filled in questionnaires, the average score was computed for questionnaires where only one item was missing. This value was then entered for unanswered questions. Missing values were not replaced from questionnaires with more than one item missing. The total score was disregarded but the scores for the individual items were used.

Results Study 1

Factor analysis

A parallel factor analysis was performed using the 10-item pool. We extracted 2 components, using optimal implementation of Parallel Analysis (PA) for determining the number of dimensions. We used Principal Components Analysis and Direct Oblimin rotation. The data were not normally distributed. Analysis of the Mardia's (1970) multivariate asymmetry found a significant kurtosis (Coefficient 135.08, statistic 5.42, $p < 0.001$). Therefore, the polychoric analysis was run. No items were removed as they met all of the following criteria that were set for each item: communality > 0.3 , factor loading on at least one factor > 0.5 , no complex factor loadings; indicated by loading > 0.4 on more than a single factor. Fifteen (10.8%) participants had missing data and they were excluded from this analysis, resulting in $n=124$ participants with BDD for this analysis. The Kaiser-Meyer-Olkin (KMO)

measure verified sampling adequacy for the analysis. The Bartlett's test of sphericity ($\chi^2 = 558.1$, $df = 45$, $p < .00001$) indicated that correlations between items were sufficiently large for factor analysis, as the Kaiser-Meyer-Olkin (KMO) test = 0.87 (good), and the determinant was 0.009. Two factors were extracted, and the factor loadings after rotation are reported in Table 1. The first factor is characterised by avoidance (accounting for 37.5% of variance), and the second factor by threat monitoring (accounting for 22.8% of the variance).

Internal Consistency

Reliability analysis using the BDD group data showed the AAI had good internal consistency with Cronbachs $\alpha = .86$.

Convergent validity

The AAI was significantly correlated with the BDD-YBOCS with a moderate coefficient indicating a strong relationship between AAI and symptoms of BDD (Table 2). There was also a significant moderate correlation with the PHQ-9 and the GAD-7 showing that higher scores on the AAI are associated with greater symptoms of generalized anxiety and depression. The AAI shared a small but significant correlation with the BABS and a moderate negative correlation with the quality of life affected by body image (BIQLI). Higher scores on the AAI are therefore associated with more negative appearance beliefs about oneself, and with greater negative impact of body image on a person's quality of life.

Sensitivity to treatment

AAI scores were assessed during treatment in 12 participants with BDD on a weekly basis to determine sensitivity to change. Therapists followed a protocol for cognitive behaviour therapy as described by Veale and Neziroglu (2010) with a maximum of 16 sessions and 3 follow-ups (at one month, three months, and one year). Figure 1 shows change on the AAI in the sample receiving CBT for their BDD. The AAI scores reduced significantly during treatment from a Median of 26.00 (IQR 9.50) to 10.50 (7.50) over 16 weeks ($Z = -2.81, p < .01$). There was a corresponding significant reduction in the scores of BDD-YBOCS from a Median of 38.50 (IQR 11.25) to 23.00 (12.00) ($Z = -2.666, p < .01$). Scores on individual items on the AAI before treatment were compared with scores at the final treatment session 16.

Method Study 2

In study 2, we validated the AAI in a community sample.

Participants

We recruited 108 participants from the community who were characterised by reporting high appearance concerns. The median age of the group was 28.5 (IQR = 14). This group consisted of 26 (24.1%) males and 82 (75.9%) females. Sixty-five were single; 34 (31.5%) were married or cohabiting; 9 (8.2%) were divorced. There were 10 unemployed or on long-term sick leave, 73 employed or self-employed, 3 retired, 19 students, and 2 homemakers and 1 missing.

Procedure

Participants were identified and recruited by an e-mail sent to 1833 volunteers (aged above 18) on the Mind Search database at the Institute of Psychiatry, King's College London. This database contains details of individuals in the local community who have volunteered to participate in psychological or psychiatric research. To determine high appearance concerns, recipients were asked if they wanted to improve one or more features in their appearance and were motivated to have a cosmetic procedure (for example, if the person was planning a cosmetic procedure or waiting until they could afford a procedure). Select-Survey ASP (TM) version 8.1.1 was used to create a web-based version of the questionnaires. They were invited to repeat the AAI one week later. The first questionnaire was completed by 108 participants (5.9% of those who were invited to take part). Sixty-seven participants (62% of total sample) repeated the AAI questionnaire one week later to determine test-retest reliability of the AAI. Participants received a £10 high street gift voucher in thanks for their time.

Materials

The participants completed the following questionnaires:

1. Appearance Anxiety Inventory (described above).
2. Cosmetic Procedure Screening (COPS) questionnaire (Veale et al, 2011).

The COPS questionnaire comprises 9 items. Items are scored from 0 (least impaired) to 8 (most impaired). It is designed to measure severity of BDD symptoms and identify people with BDD. The total scores range from 0 to 72 with higher scores reflecting increased symptoms of BDD. A score above 40 reflects increased likelihood of a diagnosis of BDD. The scale is free to download from

www.kcl.ac.uk/cadat/ under “Research”, “Questionnaires”, and “Body Image Questionnaires”. The Cronbach’s alpha in our community group (N= 108) was .86.

3. Hospital Anxiety and Depression Scale (HADS) (Zigmond & Snaith, 1983).

The HADS is a 14-item self-report instrument to screen for severity of symptoms of depression and anxiety over the past week. There are two seven-item sub-scales for anxiety and depression, with a range of scores from 0-21 on each sub scale. The standard cut offs to indicate clinical levels of depression or anxiety are 11. Cronbach’s alpha has been high in a number of validation studies (Bjelland, Dahl, Haug, & Neckelmann, 2002).

4. The Multidimensional Body Self Relations Questionnaire (MBSRQ) (Brown, Cash, & Mikulka, 1990). Two sub-scales of the MBSRQ were used: The Body Areas Satisfaction Subscale (BASS) and Appearance Evaluation (AE) subscales. The Body Areas Satisfaction subscale consists of nine items that assess dissatisfaction and satisfaction with a variety of different body areas or features related to the face or hair. The Appearance Evaluation Subscale contains 7 items to describe general body image evaluation. Higher scores on these subscales indicates greater appearance satisfaction and more positive body image evaluation.

Statistical Analysis

Horn’s Parallel Factor Analysis (Horn, 1965) was used to examine the factorial validity as described in the Statistical Analyses for Study 1. The reliability of the AAI was evaluated using Cronbach’s alpha. Total scores of the AAI were correlated using

Spearman's rho with the COPS, HAD and the MBSRQ to test convergent validity.

We used the same strategy as described in Study 1 for missing data.

Results Study 2

Factor analysis

To replicate the factor analysis of the AAI in the community group, 2 components were identified for extraction using Principal Components Analysis and Direct Oblimin rotation described in Study 1. Analysis of the Mardia's (1970) multivariate asymmetry found a significant kurtosis (coefficient 149.40; Statistic 9.860; $P = 0.0000$). Therefore the polychoric analysis was run. The Kaiser-Meyer-Olkin (KMO) measure verified sampling adequacy for the analysis and Bartlett's test of sphericity ($\chi^2 = 535.7$, $df = 45$, $p < .0001$) indicated that correlations between items were sufficiently large for factor analysis, as the Kaiser-Meyer-Olkin (KMO) test = 0.92 (very good), and the determinant of the matrix was 0.005. Two factors were extracted but it was unstable as only two items loaded on Factor 2 (Threat Monitoring). These were item 2 which had a loading of 0.89 and Item 9 that had a complex loading of 0.44 on factor 2 and of 0.40 on Factor 1. Therefore one-factor solution was attempted with the same procedure as described in the factor analysis in Study 1. The Kaiser-Meyer-Olkin (KMO) test = 0.9212 and the determinant of the matrix = 0.0055. The Bartlett's test of sphericity was 535.7 ($df = 45$; $P < 0.001$). All the items loaded between 0.37 (item 1) and 0.83 (item 3) and this accounted for 54% of the variance in scores on the AAI.

Internal Consistency

Reliability analysis showed that the AAI had high internal consistency in the community group with Cronbach's $\alpha = .91$.

Test-retest Reliability

The AAI had good test–retest reliability after 1 week with an intra-class correlation of 0.87 ($p < .001$). The first administration had a mean of 15.45, SD= 8.68, and range 1 to 38. The second administration had a mean of 16.03, SD= 8.74, and range 2 to 35.

Convergent validity

The AAI correlated moderately with the HAD depression subscale and anxiety subscale (Table 3). The AAI correlated highly with the COPS indicating that the process measure is correlated with diagnostic symptoms of BDD. The AAI also correlated moderately negatively with the MBSRQ (BASS subscale) $-.669$, $p < .001$) and the MBSRQ (AE subscale) ($\rho = -.692$, $p < .001$) indicating that higher scores on the AAI are associated with lower body satisfaction and positive appearance evaluation respectively.

Group differences

Participants in the BDD group (Mdn=27.0, IQR=12.0) scored significantly higher than the community sample (Mdn=13.0, IQR=13.5) on the AAI ($U = 3020.50$, $n = 247$, $Z = -8.06$, $p < .001$, $r = .51$).

Discussion

The current study set out to develop and validate the psychometric properties of a new process measure of appearance anxiety in BDD, to assist in research and clinical

settings. Analysis of the AAI in a clinical and community sample established the inventory's reliability, validity, and factor structure. The AAI demonstrated good internal consistency, which shows that the scale was measuring a coherent general construct.

The AAI was shown to have adequate temporal stability as measured over one week test-retest period in the community group. Evidence of the validity of the AAI in measuring symptoms of BDD came from its significant and strong associations with BDD symptoms measured by the YBOCS and COPS and depression and anxiety. The AAI also demonstrated its sensitivity to change during treatment. Furthermore total AAI scores were significantly higher in individuals with BDD compared to a non-clinical sample. Individuals treated for BDD showed significant reductions in AAI scores similar to the YBOCS-BDD during CBT.

The AAI consisted of two latent factors in the BDD group. The first of these represented avoidance and camouflaging; the second factor was comprised of threat monitoring. These factors represent the classic way of responding by either approach or avoidance of a potential threat. The items on the avoidance factor refer to avoiding cues that might trigger negative evaluation by self or others, trying to camouflage the self or avoiding any reminders of appearance. The item on comparing is partly a complex factor as it loads on both avoidance (0.59) and to a lesser extent on threat monitoring (0.29). This may be because the process of comparing a feature consists of social ranking. Individuals might fear rejection or humiliation and after ranking themselves as inferior (threat monitoring), they are likely to enact submissive behaviour and social avoidance. The “threat monitoring” component includes items

on checking in an attempt to verify exactly how one looks. This includes being excessively self-focussed and checking the internal image against that seen in a mirror or reflective surface. Ruminating and questioning others may be in response to a memory of an image in an attempt to verify exactly how one looks and problem solve. Individuals with BDD might therefore switch from a classic conflict between approach and avoidance. Alternatively some people with BDD may opt mostly for one strategy or the other. Inevitably these processes have the unintended consequence of strengthening the behaviour, increasing preoccupation and distress, and maintaining the symptoms of BDD. Therefore the main strategy of CBT for BDD after an engagement is (i) to alter the context or meaning of a distorted image through imagery re-scripting or other interventions, (ii) to stop all the cognitive and behavioural responses measured by the AAI that maintain the preoccupation and distress of BDD and a poor quality of life.

Two factors on the AAI were not replicated in Study 2 in the community group with high appearance concerns. However the factors may be less distinct in a non-clinical group because approach or avoidance may be less differentiated at lower frequencies. This may lead to one factor based on severity characterized by threat monitoring leading to avoidance and vice versa. At higher frequencies, the same processes may be more distinct.

Implications

The immediate implication from this study is that the AAI can be used as a potential measure for use in clinical and research settings for assessment over the course of a treatment intervention. The AAI may also be used to assess the outcome of treatments

in case studies, and in controlled and uncontrolled trials with larger samples. The AAI is brief, and so can be used frequently to assess change in both clinical and research settings. For example it may be helpful to determine which of the two factors to target first; whether some people adopt a strategy of mainly avoidance or threat monitoring and whether this predicts outcome or whether other psychological therapies or medication may be associated with a different pattern of response. It may also be of interest to validate the scale in other body image disorders such as eating disorders or people who have difficulty adjusting to a disfigurement to determine if there are differences in coping styles.

Limitations and Future Directions

There are limitations in the current research. We are not able to fully describe the comorbidity in our sample of people with BDD, although clinicians determined that BDD was their primary diagnosis according to DSMIV prior to treatment. Other limitations include the fact that the AAI does not focus on (i) the experience of body image or felt impression being distorted (ii) meta-cognitive beliefs about the image – for example, the degree to which a person believes that their image is an accurate representation of their appearance and in context; and (iii) beliefs and assumptions about being ugly or defective. These constructs would also need to be measured weekly in any mediating study of cognitive processes on symptoms.

Replication of the current results in a new sample of individuals with BDD or another body image disorder (such as bulimia nervosa or a disfigurement) would add strength to the findings, particularly the latent structure of the symptoms found in the current sample. Future research into the AAI could also examine the sensitivity of the

measure to treatment effects in a larger sample (and effect of other therapies or SSRI medication which do not target the processes identified), as well as the ability to distinguish between those who respond to treatment and those who do not. As discussed above, the AAI is designed to assist in determining if a mediator changes before measures of distress and quality of life.

Conclusion

The current study has validated a brief self-report scale that can be used by clinicians and researchers to assess the cognitive processes and behaviours that maintain symptoms of BDD. The AAI can be used as a tool for treatment planning and outcome measurement.

Acknowledgements

The National Institute supports David Veale, Nell Ellison and Ana Costa for Health Research (NIHR) Specialist Biomedical Research Centre for Mental Health award to the South London and Maudsley NHS Foundation Trust and the Institute of Psychiatry, King's College London. We should like to thank Caroline Easter and Alana Fogarty for their assistance in data entry and preparation.

References

- Bjelland, I., Dahl, A. A., Haug, T. T., & Neckelmann, D. (2002). The validity of the Hospital Anxiety and Depression Scale: An updated literature review. *Journal of Psychosomatic Research*, 52(2), 69-77.
- Brown, T. A., Cash, T. F., & Mikulka, P. J. (1990). Attitudinal body-image assessment: factor analysis of the Body-Self Relations Questionnaire. *Journal of Personality Assessment*, 55(1-2), 135-144.
- Cash, T. F., & Fleming, E. C. (2002). The impact of body-image experiences: Development of the Body Image Quality of Life Inventory. *International Journal of Eating Disorders*, 31(4), 455-460.
- Cash, T. F., Phillips, K. A., Santos, M. T., & Hrabosky, J. I. (2004). Measuring "negative body image": Validation of the Body Image Disturbance Questionnaire in a non-clinical population. *Body Image*, 1(4), 363-372.
- Eisen, J. L., Phillips, K. A., Baer, L., Beer, D. A., Atala, K. D., & Rasmussen, S. A. (1998). The Brown Assessment of Beliefs Scale: reliability and validity. *American Journal of Psychiatry*, 155(1), 102-108.
- Harvey, A., Watkins, E., Mansell, W., & Shafran, R. (2004). Cognitive behavioural processes across psychological disorders: A trans-diagnostic approach to research and treatment. Oxford: University Press
- Hayes, S. C., Villate, M., Levin, M., & Hildebrandt, M. (2011). Open, aware, and active: Contextual approaches as an emerging trend in the behavioral and cognitive therapies. *Annual Review of Clinical Psychology*, 7, 141-168.
- Hofman, S. G. (2008). Common misconceptions about cognitive mediation of treatment change: A commentary to Longmore and Worrell (2007). *Clinical Psychology Review*, 28, 67-70.

- Horn, J. (1965). A rationale and test for the number of factors in factor analysis. *Psychometrika*, 30(2), 179-185.
- Hrabosky, J. I., Cash, T. F., Veale, D., Neziroglu, F., Soll, E. A., Garner, D. M., Phillips, K. A. (2009). Multidimensional body image comparisons among patients with eating disorders, body dysmorphic disorder, and clinical controls: A multisite study. *Body Image*, 6(3), 155-163.
- Kraemer, H. C., Wilson, G. T., Fairburn, C. G., & Agras, W. S. (2002). Mediators and Moderators of Treatment Effects in Randomized Clinical Trials. *Archives of General Psychiatry*, 59(10), 877-883. doi: 10.1001/archpsyc.59.10.877
- Kroenke, K., & Spitzer, R. L. (2002). The PHQ-9: A new depression diagnostic and severity measure. *Psychiatric Annals*, 32(9), 509-515.
- Longmore, R. J., & Worrell, M. (2007). Do we need to challenge thoughts in cognitive behavior therapy? *Clin Psychol Rev*, 27(2), 173-187.
- Lorenzo-Seva, U., & Ferrando, P. J. (2006). FACTOR: A computer programme to fit the exploratory factor analysis model. *Behavioral Research Methods, Instruments and Computers*, 38, 88-91.
- Mancuso, S. G., Knoesen, N. P., & Castle, D. J. (2010). The Dysmorphic Concern Questionnaire: A screening measure for body dysmorphic disorder. *The Australian and New Zealand Journal of Psychiatry*, 44(6), 535-542.
- Mardia, K. V. (1970). Measures of multivariate skewness and kurtosis with applications. *Biometrika*, 57(3), 519-530.
- Neziroglu, F., Khemlani-Parel, S., & Veale, D. (2008). Social learning theory and cognitive behavioural models of body dysmorphic disorder. *Body Image*, 5(1), 28-38.
- Phillips, K. A., Hollander, E., Rasmussen, S. A., Aronowitz, B. R., DeCaria, C., & Goodman, W. K. (1997). A severity rating scale for body dysmorphic disorder:

development, reliability, and validity of a modified version of the Yale-Brown Obsessive Compulsive Scale. *Psychopharmacology Bulletin*, 33(1), 17-22.

Rosen, J. C., Reiter, J., & Orosan, P. (1995). Cognitive-behavioral body image therapy for body dysmorphic disorder [published erratum appears in J Consult Clin Psychol 1995 Jun;63(3):437]. *Journal of Consulting & Clinical Psychology*, 63(2), 263-269.

Spitzer, R. L., Kroenke, K., Williams, J. B. W., & Lowe, B. (2006). A brief measure for assessing generalized anxiety disorder - The GAD-7. *Archives of Internal Medicine*, 166(10), 1092-1097. doi: 10.1001/archinte.166.10.1092

Veale, D. (2004). Advances in a cognitive behavioural model of body dysmorphic disorder. *Body Image*, 1, 113-125.

Veale, D., Gournay, K., Dryden, W., Boocock, A., Shah, F., Willson, R., Walburn, J. (1996) Body Dysmorphic Disorder: a cognitive behavioural model and a pilot randomised controlled trial. *Behaviour, Research and Therapy*, 34, 9:717-729

Veale, D., Ellison, N., Werner, T. G., Dodhia, R., Serafty, M., & Clarke, A. (2011). Development of a cosmetic procedure screening questionnaire (COPS) for body dysmorphic disorder. *Journal of Plastic, Reconstructive & Aesthetic Surgery* 65(4), 530-532. doi:10.1016/j.bjps.2011.09.007

Veale, D., & Neziroglu, F. (2010). *Body Dysmorphic Disorder: a treatment manual*. Chichester: John Wiley & Sons.

Wilson, P., & Cooper, C. (2008). Finding the Magic Number. *The Psychologist*, 21(10), 866-867.

Worrell, M., & Longmore, R. J. (2008). Challenging Hofmann's negative thoughts: A rebuttal. *Clinical Psychology Review*, 28, 71-74.

Zigmond, A., & Snaith, R. P. (1983). The hospital depression and anxiety scale. *Acta Psychiatrica Scandinavica*, 67, 361-370.

Zwick, W. R., & Velicer, W. F. (1986). Comparison of five rules for determining the number of components to retain. *Psychological Bulletin*, 99, 432-442.

Table 1

Factor Loadings for the Final Two-Factor Solution in Participants with BDD (Study 1)

Item	Factor loadings	
	Factor 1	Factor 2
I check my appearance (e.g. in mirrors, by touching with my fingers or by taking photos of myself)	0.29	0.59*
I compare aspects of my appearance to others	0.58*	0.22
I avoid situations or people because of my appearance	0.84*	-0.09
I think about how to camouflage or alter my appearance	0.81*	0.10
I avoid reflective surfaces, photos or videos of myself	0.64*	-0.13
I try to camouflage or alter aspects of my appearance	0.85*	-0.01
I brood about past events or reasons to explain why I look the way I do	0.16	0.67*
I am focussed on how I feel I look rather than on my surroundings	0.28	0.66*
I discuss my appearance with others or question them about it	-0.19	0.87*
I try to prevent people from seeing aspects of my appearance within particular situations (e.g. by changing my posture, avoiding bright lights)	0.71*	0.09

Note * $p < .05$

Table 2

Correlations between AAI and additional measures (Study 1)

	YBOCS-BDD	BABS	PHQ9	GAD	BIQLI
AAI	0.55 **	0.31 *	0.58**	0.55**	-0.54**

** P < 0.001, * P< 0.05

Table 3

Correlations between AAI and additional measures (Study 2)

	Depression	Anxiety	COPS	BASS	AE
AAI	0.65 **	0.63**	0.67**	0.67**	-0.69**

** P < 0.001